

Personalizing medical decision making

Recent advances in prediction model research





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Research of statistical methods

- Risk prediction
- Evidence synthesis







My talk today

- What is prediction?
- Recent advances in Machine Learning
- Recent advances in Penalization
- Recent advances in Evidence Synthesis
- Recent advances in Treatment effect modelling
- Next Steps





Data sharing for Infectious Diseases

A federated cloud for OMICS and clinical data

Re CoD ID

Reconciliation of Cohort Data for Infectious Diseases

CIHR The ReCoDD project has received funding from European Union's Horizon 2020 Research and Innovation Programme under Grant Agreement No.825746 IRSC and is supported by the Canadian Institutes of Health Research, Institute of Genetics (DHR-IG) under Grant Agreement Nº 01886-000





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Background





Estimate something that is yet unknown

- Presence of a certain disease (diagnosis)
- Future occurrence of a particular event (prognosis)







Calculate the absolute risk (probability) for distinct individuals

Why?

- Identify high-risk individuals
- Identify absolute treatment effect
- Target decision making to individuals







Calculate the absolute risk (probability) for distinct individuals

How?

Combine information from multiple predictors

- Subject characteristics (e.g. age, gender)
- History and physical examination results (e.g. blood pressure)
- Imaging results
- (Bio)markers (e.g. coronary plaque)





Calculate the absolute risk (probability) for distinct individuals

| iender: | FEMALE | |
|----------------------------------|--------|--|
| ge: | 40 | |
| moker: | NO | |
| ystolic blood ressure (mmHg): | 120 | |
| liabetes: | NO | |
| holesterol mg/dl): | 200 | |
| Calcı | ulate | |

| RES | ULTS | |
|---|-----------------------|--|
| 10 year risk of CV eve CV risk is Low. | ent: | |
| More recommendations > | <10% | |
| Input data | | |
| Gender: Female | Age: 40 | |
| Cholesterol (mg/dl): 200 | | |
| Systolic blood pr 120 | essure (mmHg): | |
| Smoker: No D | iabetes: No | |
| What would happen if | | |
| Smoker: | NO | |
| Systolic blood pressure (mmHg): | 120 | |
| Cholesterol | | |
| CV RISK BODY MASS INDEX | RECOMMENDATIONS ALARM | |



Develop a multivariable statistical model

- Need for patient data from large cohort studies
- Many strategies available (Regression, decision trees, neural networks)



What is a good model?





What is a good model?

Reproducitbility versus Transportability

- Performance in same population*
- Evaluated with:
 - Internal validation (resampling methods using random-split)
 - External validation (same population)

- Performance in a different but related population*
- Evaluated with:
 - External validation (different population)
 - Resampling methods with non-random split



[*] Debray, T., Vergouwe, Y., Koffijberg, H., Nieboer, D., Steyerberg, E. and Moons, K. (2015). A new framework to enhance the interpretation of external validation studies of clinical prediction models. *Journal of Clinical Epidemiology*, 68(3), pp.279-289.

Current limitations

Many prediction models perform poorly, do not affect clinical practice, or do not improve patient outcomes

- Small & poor quality studies
- Limited variation in studied patients, settings or populations
- Lack of validity and effectiveness assessments

"All models are wrong, but some are useful" – George Box



Improving prediction models

- Machine Learning
- Penalization
- Evidence synthesis
- Treatment effect modelling











https://medium.com/analytics-vidhya/fundamental-omachine-learning-ada28afa1bd3



"There are two cultures in the use of statistical modeling to reach conclusions from data. One assumes that the data are generated by a given stochastic data model. The other uses algorithmic models and treats the data mechanism as unknown." – **Leo Breiman**



Statistical Science, Vol. 16, No. 3 (2001), pp. 199-215

- Strong focus on prediction and classification
- Combination of data-driven algorithms
 - Nearest Neighbour
 - Recursive Partitioning
 - Neural Network
 - Support Vector Machine
- Avoidance of modeling assumptions (e.g. additivity, linearity), resulting in high flexibility





Data available for prediction:

- Imaging (e.g. CT scan, MRI)
- Text (e.g. medical records)
- High-throughput data (e.g. wearables)
- High-dimensional laboratory data
- Clinical epidemiological data





Major contributions

- Image recognition
- Analysis of unstructured data
- Problems with high signal:noise ratio



Major challenges

- Severe overfitting in "small" samples
- Very limited gains in the analysis of large (structured) epidemiological datasets
- Not designed for causal inference



Modern modelling techniques are data hungry: a simulation study for predicting dichotomous endpoints

Tjeerd van der Ploeg^{1,3*}, Peter C Austin² and Ewout W Steyerberg³

Modern modeling techniques had limited external validity in predicting mortality from traumatic brain injury

Tjeerd van der Ploeg^{a,b,c,*}, Daan Nieboer^c, Ewout W. Steyerberg^c
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Accepted 5 March 2016; Published online 14 March 2016

Logistic regression has similar performance to optimised machine learning algorithms in a clinical setting: application to the discrimination between type 1 and type 2 diabetes in young adults

Anita L. Lynam¹, John M. Dennis¹, Katharine R. Owen^{2,3}, Richard A. Oram¹, Angus G. Jones¹, Beverley M. Shields¹ and Lauric A. Ferrat^{1*}

A systematic review shows no performance benefit of machine learning over logistic regression for clinical prediction models Evangelia Christodoulou^a, Jie Ma^b, Gary S. Collins^{b,c}, Ewout W. Steyerberg^d, Jan Y. Verbakel^{a,e,f}, Ben Van Calster^{a,d,*}



Machine Learning may not (yet) be suitable for prediction of absolute treatment effects in routine care settings



Penalization

Improved prediction of time-to-event outcomes



The need for penalization

Many prediction models are prone to overfitting

- Noise is (partially) interpreted as signal
- Inaccurate predictions for new patients from the target population
 - Predicted risk is too high for high-risk patients
 - Predicted risk is too low for low-risk patients
- Estimates of out-of-sample performance are over-optimistic





The need for penalization

How to avoid overfitting?

- Regularize model complexity (e.g. via assumptions)
- Shrink poorly calibrated predictions towards the average risk
- Constrain the magnitude of regression coefficients
- Include a penalty term in the log-likelihood
- Examples: LASSO, Ridge, Elastic Net, etc.



Overfitting: an example



Observed proportions versus average predicted risk of the event



Pavlou M, et al. BMJ. 2015;351:h3868.

What about prediction of time-to-event outcomes?

- Need for parametric survival models
- Need for flexible baseline hazard (BH)
- Need for penalization

| Model type | Parametric BH | Flexible BH | Penalization |
|----------------|---|---|--------------|
| Cox | Х | Image: A set of the set of the | × |
| Weibull | × | Х | × |
| Royston-Parmar | Image: A set of the set of the | × | X |

Research by drs. Jeroen Hoogland

- Combine the benefits of the Royston Parmar log cumulative hazards model and penalized maximum likelihood estimation
- Implement an elastic net penalty for the RP model
- Facilitate estimation of non-proportional hazards and other interaction terms







- The log cumulative hazard is modeled as a linear additive combination
- All terms are differentiable w.r.t. (log) time
- Thus, the log-likelihood is available in closed form
- Penalty

$$P_{\omega}(\boldsymbol{\theta}) = \sum_{i=1}^{d} \omega_i \lambda_{1i} |\theta_i| + (1 - \omega_i) \frac{1}{2} \lambda_{2i} \theta_i^2$$

- The size of the penalty can be modified per parameter (lambda)
- The mixture between ridge and lasso can be modified per parameter (omega)

- Full gradient ascent algorithm (based on lasso Cox PH)
- Step size depends on ratio l'_{pen} / l"_{pen}
 - First derivative of the penalized log-likelihood l'pen
 - Second derivative of the penalized log-likelihood I"pen
- Respects discontinuities in the gradient for parameters subject to an absolute value penalty
- When close to the optimum, switches to Newton-Raphson
- Hyper-parameter tuning using out-of-sample log-likelihood

Simulation study

Data simulated from a Weibull mixture with non-proportional hazards

- 20 MVN covariates with mutual correlation 0.25
 - 12 noise variables
 - 8 variables with beta = 0.25
 - 1 (independent) treatment variable with beta = -0.5
- Survival times were right-censored (administrative)
- Event rate ~ 0.75
- 500 patients available for model development
- 5000 patients for model evaluation

Simulation study results

Error in predicted survival (q .1, .25, .5, .75, .9)





Simulation study results

Error in predicted individual treatment effect





Main findings

- The Royston Parmar log cumulative hazards model is very flexible
- Model complexity often needs to be tuned to the data at hand
- Regularization provides a means to do so

Limitations

- algorithm is sensitive to starting values
- As of yet, it starts from ML and PH based initial values
- Therefore, is does not scale well in case of strongly non-PH models with >> p



Overfitting – a problem solved?

Findings from a recent simulation study

- Despite improved performance on average, shrinkage often worked poorly in individual datasets, in particular when it was most needed.
- Shrinkage methods do not solve problems associated with small sample size or low number of events per variable



Article

Regression shrinkage methods for clinical prediction models do not guarantee improved performance: Simulation study



Statistical Methods in Medical Research 0(0) 1–13 © The Author(s) 2020 Article reuse guidelines: sagepub.com/journals-permissions DOI: 10.1177/0962280220921415 journals.sagepub.com/home/smm **SAGE**
Overfitting – a problem solved?

- Traditional penalization methods help to improve performance when the model is applied to new patients from the same target population (i.e. reproducibility)
- Penalization does not aim to improve the model's performance when applied across different (but related) settings and populations (i.e. transportability)

More advanced methods are needed to quantify and improve the generalizability of prediction models



Evidence synthesis

Improving predictions across different settings and populations



Evidence synthesis in prognosis research

Synthesis of prognosis studies may help

- To identify promising markers
- To identify promising prediction models
- To improve the accuracy of prediction models



Evidence synthesis in prognosis research

- Meta-analysis of published aggregate data (AD)
 - Summarize prediction model performance
 - Summarize risk factor-outcome associations
- Meta-analysis of individual participant data (IPD)
 - Develop & validate prediction models
 - Identify prognostic factors
 - Identify predictors of treatment effect
- Meta-analysis of IPD and AD



Meta-analysis of published AD

Research Methods & Reporting

A guide to systematic review and meta-analysis of prognostic factor studies

BMJ 2019 ; 364 doi: https://doi.org/10.1136/bmj.k4597 (Published 30 January 2019) Cite this as: *BMJ* 2019;364:k4597

Research Methods & Reporting

A guide to systematic review and meta-analysis of prediction model performance

the**bm**

BMJ 2017 ; 356 doi: https://doi.org/10.1136/bmj.i6460 (Published 05 January 2017) Cite this as: *BMJ* 2017;356:i6460



Meta-analysis of published AD

Guidance for systematic reviews (research by dr. Damen)

- Defining the review question (PICOTS)
- Defining the search strategy
- Quantitative data extraction
- Quality appraisal (PROBAST, QUIPS)
- Meta-analysis (metamisc R package)
- Investigating between-study heterogeneity
- Interpretation (GRADE)
- Reporting (guidelines: REMARK, PRISMA, TRIPOD)









Meta-analysis of diagnostic and prognostic modelling studies



https://CRAN.R-project.org/package=metamisc



Meta-analysis of published AD

Recent reviews to summarize prediction model performance

- Breast cancer (Meads *et al*; Breast Cancer Res. Treat. 2012)
- Perioperative Mortality (Sullivan *et al*; Am. J. Cardiol. 2016)
- Cardiovascular disease (Damen *et al*; BMC Med 2017)
- Colorectal cancer (Hu et al; Surg Oncol 2019)
- Chronic lymphocytic leukemia (Molica *et al*; Leukemia 2020)
- •



Meta-analysis of IPD

Data increasingly available for thousands or even millions of patients from multiple practices, hospitals, or countries.

- Meta-analysis of individual participant data (IPD) from multiple studies
- Analyses of databases and registry data containing e-health records





Meta-analysis of IPD

Main opportunities

- Increase total sample size
- Increase available case-mix variability
- Ability to standardize analysis methods across IPD sets
- Ability to investigate more complex associations
- Ability to evaluate generalizability of the model across different settings and populations



External validation using IPD-MA

Validation of QRISK 2 in 364 UK practices



Riley RD, et al. BMJ. 2016;353:i3140.

Model development using IPD-MA

Internal-external cross-validation



Debray TPA, et al. Stat Med. 2013 Aug 15;32(18):3158–80. Steyerberg EW, Harrell FE. J Clin Epidemiol. 2015 Apr 18;69:245–7.

Development of ENCALS

Prognosis of amyotrophic lateral disease

- 14 cohort studies (specialized ALS centres)
 - N = 190 to 1,936 per study (total N = 11,475)
 - Median follow-up: 97.5 months
 - Composite endpoint

(Non-invasive ventilation for more than 23h/day, or death)



Development of ENCALS

c statistic (95% CI)

| Utrecht, Netherlands | | 0.79 (0.77 to 0.81) |
|------------------------|------------------|-----------------------|
| Dublin, Ireland | | 0.78 (0.76 to 0.80) |
| Torino, Italy | | 0.77 (0.75 to 0.79) |
| Sheffield, UK | | 0.78 (0.76 to 0.80) |
| London, UK | | 0.82 (0.79 to 0.84) |
| Oxford, UK | _ _ | 0.78 (0.75 to 0.81) |
| Leuven, Belgium | | 0.77 (0.75 to 0.80) |
| Lisbon, Portugal | _ | 0.77 (0.74 to 0.80) |
| Hannover, Germany | | 0.74 (0.71 to 0.77) |
| Ulm, Germany | | - 0.83 (0.78 to 0.88) |
| Jena, Germany | _ | 0.80 (0.75 to 0.85) |
| St Gallen, Switzerland | | 0.80 (0.74 to 0.86) |
| Tours, France | _ | 0.76 (0.71 to 0.81) |
| Limoges, France | _ | 0.80 (0.73 to 0.86) |
| Meta-analysis | _ | 0.78 (0.77 to 0.80) |
| 0.7 | 0 0.75 0.80 0.85 | 95% PI 0.74 to 0.82 |

Validation cohort

THE LANCET Neurology



Volume 17, Issue 5, May 2018, Pages 423-433

Articles

Prognosis for patients with amyotrophic lateral sclerosis: development and validation of a personalised prediction model Henk-Jan Westeneng MD^a, Thomas P A Debray PhD^{b, o}, Anne E Visser MD^a, Ruben P A van Eijk MD^a, James P K Rooney MSc ^d, Andrea Calvo MD^e, Sarah Martin BSc ^r, Prof Christopher J McDermott PhD 9, Alexander G Thompson BMBCh^h, Susana Pinto PhD¹, Xenia Kobeleva MD¹, Angela Rosenbohm MD^k, Beatrice Stubendorff PhD¹, Helma Sommer^m, Bas M Middelkoop^a, Annelot M Dekker MD^a, Joke J F A van Vugt PhD^a, Wouter van Rheenen MD^a, ... Prof Leonard H van den Berg MD^a, S

| Performance | Criteria | Prob. of "good" performance | Joint probability | |
|--------------------------|-----------------|--------------------------------|----------------------|--|
| C-statistic | > 0.70 | 100% | 00.20/ | |
| Calibration slope | 0.80 to 1.20 | 97.1% | 98.3% | |
| Calibration-in-the-large | -0.587 to 0.587 | 85.5% | | |

Developing generalizable prediction models

Stepwise estimation procedure (research by dr. de Jong)

- Fitting of a pre-specified GLM in each study
- Evaluation of performance using IECV
- Loss = f(overall performance in hold-out studies, between-study variation)
- Expand (or reduce) model until the overall loss no longer decreases
- Implementation in "metamisc"





Developing generalizable prediction models

Further extensions

- Methods to adjust for measurement error in IPD-MA
- Methods to disentangle case-mix variation from invalid predictor effects
- Methods to account for missing participant-level data in IPD-MA

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Treatment effect modelling

Improving predictions of absolute treatment effect





Individualized absolute treatment effects provide a natural starting point to engage in shared decision making

Requirements

- Move to the absolute risk scale
- Adjust for individual patient characteristics
- Consider counterfactual outcomes







Individualized absolute treatment effects provide a natural starting point to engage in shared decision making

Two important sources of information (in RCTs):

- Prognostic variables predicting outcome risk on reference treatment
- Treatment variables with potential for effect modification







An example: The SYNTAX score II

"The SYNTAX score II is a clinical tool that combines clinical variables with the anatomical SYNTAX score, providing expected 4-year mortality for both coronary artery bypass grafting (CABG) and percutaneous coronary intervention (PCI) — thus recommending either PCI only, CABG only or equipoise in treatment based on long-term mortality."

DOI: 10.21037/acs.2018.07.02







SYNTAX SCORE II 4-year mortality

SYNTAX Score II questions



5



SYNTAX Score II



Decision making -between CABG and PCI- guided by the SYNTAX Score II to be endorsed by the Heart Team.

| PCI SYNTAX Score II: PCI 4 Year Mortality: | 46.6 24.7 % Absolute treatment effect is | |
|--|--|--|
| CABG SYNTAX Score II: CABG 4 Year Mortality: | 26.8 19.4% in favor of CABG | |
| Treatment recommendation (i): | CABG or PCI | |





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Problem definition

- Randomized Clinical Trials are designed for estimating relative treatment effects (e.g. RR, OR)
- Can we use RCT data to predict more individualized absolute treatment effects?



Development of treatment effect models

Aim: To compare regression modeling methods on their ability to predict individual absolute treatment effect

Investigation of treatment effect models

Based on logistic regression

- Global: a single model for the whole population
- Partitioning: multiple simple models for partitions of the population







Global models

Absence of HTE (risk magnification)

$$logit(P(Y_i = 1 | \boldsymbol{x}_i, t_i = 0)) = \beta_0 + \boldsymbol{x}_i^{\mathsf{T}} \boldsymbol{\beta} = \eta_i \qquad \#df = p+2$$
$$\hat{\delta}_i = P(Y_i = 1 | \eta_i)) = \frac{1}{1 + e^{-(\eta_i + \beta_t)}} - \frac{1}{1 + e^{-\eta_i}}$$

Presence of HTE (individual treatment-covariate interactions)

$$logit(P(Y_i = 1 | \boldsymbol{x}_i, t_i)) = \beta_0 + t_i \beta_1 + \boldsymbol{x}_i^\top \boldsymbol{\beta}_m + t_i \boldsymbol{x}_i^\top \boldsymbol{\beta}_z \qquad \#df = 2p+2$$

Presence of HTE (interaction between treatment and baseline risk)

$$\operatorname{ogit}(P(Y_i = 1 | \eta_i, t_i)) = \gamma_0 + t_i \gamma_1 + \eta_i + t_i f(\eta_i) \qquad \#df = p + 3 + (1 + 1)$$



Partitioning models

- Model-based recursive partitioning
- Start with a simple global model $logit(P(Y_i = 1|)) = \beta_0 + t_i\beta_1$
- Form partitions \mathcal{B}_b in the space of $\mathcal{X} = X_1 \times \ldots \times X_p$ such that $\operatorname{logit}(P(Y_i = 1 | \mathcal{B}_b)) = \beta_{0b} + t_i \beta_{1b}$ holds
- Implemented as
 - 1. Variable-by-variable subgroup selection (single split)
 - 2. Single tree
 - 3. Random forest



Methods for treatment effect modelling

| Methods | Equations | Estimation |
|---|-------------|----------------------------|
| Global | | |
| Overall absolute treatment effect (Overall) | — | ML |
| Risk magnification | (2),(3) | ML, Elastic net |
| marginal treatment effect (RMm) | | |
| Risk magnification | (1),(2),(3) | ML, Elastic net |
| conditional treatment effect (RMc) | | |
| Baseline risk modifier approach | (2),(5) | ML, Elastic net |
| linear treatment interaction (BA_linear) | | |
| Full modeling (FM) | (4) | ML, Elastic net, |
| | | HGL, Boosting |
| Partitioning | | |
| Single subgroup | (6) | MOB stump |
| Single tree | (6) | MOB |
| Random forest | see [17] | pMOB |



ML = Maximum Likelihood, HGL = Hierarchical Group Lasso

Empirical example

- RCT with 1:1 allocation ratio (N = 512)
- Population: clinically diagnosed acute otitis media (AOM) in children 6 months to 5 years of age
- Intervention: amoxicillin
- Outcome: fever or ear pain was after 3 days' follow-up
- Baseline data on: treatment received, sex, presence of recurrent AOM, fever, bilateral occurrence, ear pain, presence of a runny nose, cough, tympanic membrane abnormality, and age



Empirical example





Empirical example





Simulation study

- Logistic data generating mechanism
 - 1:1 allocation ratio
 - 20% event rate
 - 6 covariates with a main effect (MVN with rho = 0.3)
- Variable simulation parameters
 - sample size 250 or 2500
 - presence/absence of average relative treatment effect
 - number and size of treatment-covariate interactions
 - absence/presence of (6) noise variables

Simulation study results (1 interaction)

Average root Mean Squared Prediction Error (rMSPE) of the predicted absolute treatment effect





Simulation study results (3 interactions)

Average root Mean Squared Prediction Error (rMSPE) of the predicted absolute treatment effect

@TPA Debray



Conclusions

- Small RCTs
 - Hard to improve beyond risk-magnification
 - However, the price to pay to allow for treatment-covariate interactions was small when using both shrinkage and selection, especially for the hierarchical group lasso (HGL)
- Large RCTs
 - Shrinkage and selection still needed
 - Allowing for all interactions was beneficial



Conclusions

- Baseline risk modifier approach, variable-by-variable subgroup, and single MOB were always outperformed
- Random forest MOB performed relatively well given the simulation settings



Next steps

- Improving development and validation in single study
 - Penalization of absolute treatment benefit
 - Machine Learning with assumptions
 - Adjusting for competing risk
 - Quantifying accuracy of absolute treatment benefit
- Evidence synthesis
 - Meta-analysis of individual participant data and published AD
 - Meta-analysis of randomized and observational studies

